Probabilistic Aspects of Fatigue

Introduction

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Fatigue Under Complex Loading: Analysis and Experiments, SAE AE6, 1977



















Deterministic – from past measurements the future position of a satellite can be predicted with reasonable accuracy

Random – from past measurements the future position of a car can only be described in terms of probability and statistical averages



variability and uncertainty is accommodated by introducing safety factors. Larger safety factors are better, but how much better and at what cost?



Reliability = 1 – P(Stress > Strength)

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 3σ contains 99.87% of the data

P(s < S) = 2.3 10⁻³

If we use 3σ on both stress and strength

 $\text{P(failure)} = \text{P(} \ \Sigma \geq \text{s} \ \cap \ \text{s} \leq \text{S} \text{)} = 5.3 \ 10^{-6} \approx 4.5 \, \sigma$

The probability of the part with the lowest strength having the highest stress is very small

For 3 variables, each at 3 σ :

P(failure) =1.2 $10^{-8} \approx 5.7 \,\sigma$



- Reduces conservatism (cost) compared to assuming the "worst case" for every design variable
- Quantifies life drivers what are the most important variables and how well are they known or controlled ?
- Quantifies risk



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Basic Probability and Statistics



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Deterministic – from past measurements the future position of a satellite can be predicted with reasonable accuracy

Random – from past measurements the future position of a car can only be described in terms of probability and statistical averages





- Discrete fixed number of outcomes
 - Colors
- Continuous may have any value in the sample space
 - Strength



- Mean or Expected Value
- Variance / Standard Deviation
- Coefficient of Variation
- Skewness
- Kurtosis
- Correlation Coefficient





Central tendency of the data

Mean =
$$\mu_x = \overline{x} = E(X) = \frac{\sum_{i=1}^{N} x_i}{N}$$

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Variance / Standard Deviation

Dispersion of the data

$$\operatorname{Var}(\mathbf{X}) = \frac{\sum_{i=1}^{N} (\mathbf{x}_{i} - \overline{\mathbf{x}})^{2}}{N}$$

Standard deviation

$$\sigma_x = \sqrt{Var(X)}$$

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$$COV = \frac{\sigma_x}{\mu_x}$$

Useful to compare different dispersions

$$\begin{array}{ll} \mu = 10 & \mu = 100 \\ \sigma = 1 & \sigma = 10 \\ \text{COV} = 0.1 & \text{COV} = 0.1 \end{array}$$

Skewness

Skewness is a measure of the asymmetry of the data around the sample mean. If skewness is negative, the data are spread out more to the left of the mean than to the right. If skewness is positive, the data are spread out more to the right. The skewness of the normal distribution (or any perfectly symmetric distribution) is zero.

Skewness(X) =
$$\frac{\sum_{i=1}^{N} (x_i - \overline{x})^3}{N\sigma^3}$$

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Kurtosis is a measure of how outlier-prone a distribution is. The kurtosis of the normal distribution is 3. Distributions that are more outlier-prone than the normal distribution have kurtosis greater than 3; distributions that are less outlier-prone have kurtosis less than 3.

Kurtosis(X) =
$$\frac{\sum_{i=1}^{N} (x_i - \overline{x})^4}{N\sigma^4}$$



A measure of the linear association between random variables

$$\sigma_{xy} = COV(X, Y) = E[(X - \mu_X)(Y - \mu_Y)]$$





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Reliability





The probability of event A occurring:

$$\label{eq:posterior} \begin{split} 0 &\leq P\bigl(A\bigr) \leq 1 \\ P\bigl(A\bigr) &= 1 \mbox{ certain } \\ P\bigl(A\bigr) &= 0 \mbox{ imposssible } \end{split}$$

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What is the probability of the structure failing?



Bar 1 or bar 2 fails

$$P(A \cup B) = P(A) + P(B) - P(A \cap B)$$

$$P(A) = 0.03$$

$$P(B) = 0.04$$

$$P(A \cap B) = P(A) \cdot P(B) = 0.0012$$

$$P(\text{ failure}) = 0.03 + 0.04 - 0.0012 = 0.0688$$

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P(no failure) = Reliability Define P(\overline{A}) probability of not A P(\overline{A}) = P(1 - A) Reliability = P($\overline{A} \cap \overline{B}$) P($\overline{A} \cap \overline{B}$) = P(\overline{A}) · P(\overline{B}) For the 2 bar structure

 $P(\overline{A} \cap \overline{B}) = 0.97 \cdot 0.96 = 0.9312$ P(failure) = 1-Reliability = 0.0688

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6 sigma is 1 in a billion (0.999999999 Reliability)

Suppose a structure has 1000 bolted joints:

 $P(\overline{A}) = (0.999999999)^{1000} = 0.9999999$ 1 in a million

3 sigma is (0.99865 Reliability)

 $P(\overline{A}) = (0.99865)^{1000} = 0.26$

74 % failures

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 $\begin{aligned}
& \mu_{\ln x} = \ln \left(\frac{\mu_x}{\sqrt{1 + \text{COV}_x^2}} \right) \\
& \sigma_{\ln x}^2 = \ln \left(1 + \text{COV}_x^2 \right) \\
& \text{COV}_x = \sqrt{\exp(\sigma_{\ln x}^2) - 1} \\
& \mu_x = \exp(\mu_{\ln x} + 0.5 \sigma_{\ln x}^2) \\
& \overline{X}_x = \exp(\mu_{\ln x}) = \frac{\mu_x}{\sqrt{1 + \text{COV}_x^2}}
\end{aligned}$





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Estimated maximum force from the distributions

Normal	893
LogNormal	1125
Gumbel	1223
Weibull	785







 Joint Probability Density

 Normal distributions

 $f_{xy}(x,y) = \frac{1}{2\pi\sigma_x \sigma_y \sqrt{1-\rho^2}}$
 $exp\left[\frac{-1}{2(1-\rho^2)}\left\{\left(\frac{x-\mu_x}{\sigma_x}\right)^2 - 2\rho\left(\frac{x-\mu_x}{\sigma_x}\right)\left(\frac{y-\mu_y}{\sigma_y}\right) + \left(\frac{y-\mu_y}{\sigma_y}\right)^2\right\}\right]$

 ρ correlation coefficient







For most durability problems, we are not interested in the "large extremes" of stress or strength. Failure is much more likely to come from moderately high stresses combined with moderately low strengths.



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Statistical Techniques



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- Normal Distributions
- LogNormal Distributions
- Monte Carlo
- Sampling
- Distribution Fitting







Linear Response Function

$$Z = \mathbf{a}_{o} + \sum_{i=1}^{n} \mathbf{a}_{i} X_{i}$$
$$X_{i} \sim N(\mu_{i}, C_{i})$$
$$\mu_{z} = \mathbf{a}_{o} + \sum_{i=1}^{n} \mathbf{a}_{i} \mu_{i}$$
$$\sigma_{z} = \sqrt{\sum_{i=1}^{n} \mathbf{a}_{i}^{2} \sigma_{i}^{2}}$$

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 $Z = S - \Sigma$

Failure will occur whenever Z <= 0

$$Z = \mu_z - z \sigma_z = 0$$
$$Z = \frac{\mu_z}{\sigma_z} = \frac{100}{28.2}$$

z = 3.54 standard deviations

 $P(failure) = 2 \times 10^{-4}$

For this case only, a safety factor of 2 means a probability of failure of 2×10^{-4} . Other situations will require different safety factors to achieve the same reliability.

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What is the expected distribution in fatigue lives?



 $Z = a_{o} \prod_{i=1}^{n} X_{i}^{a_{i}}$ a's are constant and $X_{i} \sim LN(x_{i}, C_{i})$ median $\overline{Z} = a_{o} \prod_{i=1}^{n} \overline{X}_{i}^{a_{i}}$ $COV \quad C_{Z} = \sqrt{\prod_{i=1}^{n} (1 + C_{X_{i}}^{2})^{a_{i}^{2}} - 1}$

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	∆S/2	σ'_{f}	2N _f	Percentile	Life
μ _x	250	1000	355,368	99.9	17,706,069
COV _x	0.2	0.1	4.72	99	4,566,613
				95	1,363,200
μinx	5.50	6.90	11.21	90	715,589
Х	245	995	73,676	50	73,676
σχ	50	100	1,676,831	10	7,586
ମାnx	0.198	0.100	1.774	5	3,982
				1	1,189
b =	-0.125			0.1	307

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$$\frac{K_{f}\Delta S}{2} = \sqrt{E\left(\frac{{\sigma_{f}}^{'2}}{E}\left(2N_{f}\right)^{2b} + \sigma_{f}^{'}\epsilon_{f}^{'}\left(2N_{f}\right)^{b+c}\right)}$$

Given random variables for K_f, $\Delta S, \sigma_f$ and ϵ_f Find the distribution of 2N_f

$$Z = 2N_{f} = ?$$







- 1. Generate random numbers between 1 and 6, all integers
- 2. Count the number of 3's

Let X_i = 1 if 3 0 otherwise

$$P(3) = \frac{1}{n} \sum_{i=1}^{n} X_i$$

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•Probab

=ROUNDUP(6 * RAND(), 0) =IF(A1 = 3, 1, 0) =SUM(\$B\$1:B1)/ROW(B1)

5	0	0
3	1	0.5
4	0	0.333333
4	0	0.25
5	0	0.2
6	0	0.166667
1	0	0.142857
3	1	0.25
3	1	0.333333
6	0	0.3



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Repeat many times



$$x = F_x^{-1}(RAND)$$

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EXCEL				
	σ _f	$\frac{\Delta S}{2}$	2N _f	
	893	204	134,677	
	1102	301	32,180	
	852	285	6,355	
	963	173	929,249	
	1050	283	35,565	
	1080	265	77,057	
	965	313	8,227	
	1073	213	420,456	
	1052	226	224,000	
	954	322	5,878	
	965	240	68,671	
	993	207	277,192	
	1191	368	11,967	
	831	210	59.473	

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Simulation is relatively straightforward and simple

Obtaining the necessary input data is difficult



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Sample variance: s²

$$E(\overline{X}) = \mu_{x}$$
$$E(s^{2}) = \sigma_{x}^{2}$$



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Confidence Intervals $E(\overline{X}) = \mu_x$

What is the probability that a sample \overline{X} is greater than μ_x ? 50%

$$P(L \le \mu_x \le U) = 1 - \alpha$$

There is a $1 - \alpha$ chance of selecting a sample in the interval between L and U that contains the true mean of the population



90% confidence

If we sampled a population many times to estimate the mean, 90% of the time the true population mean would lie between the computed upper and lower limit.



Confidence Interval - mean

For a normal distribution:

Lower limit of $\boldsymbol{\mu}$

$$\overline{X} - t_{\alpha,n-1} \frac{s_x}{\sqrt{n}} \leq \mu_x$$

Upper limit of $\boldsymbol{\mu}$

$$\overline{X} + t_{\alpha,n-1} \frac{s_x}{\sqrt{n}} \geq \mu_x$$

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For a normal distribution:

Upper limit of σ

$$\frac{(n\!-\!1){s_x}^2}{\chi^2_{1\!-\!\alpha,n\!-\!1}} \le {\sigma_x}^2$$

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Chi-square

any univariate distribution

Snedecor, George W. and Cochran, William G. (1989), Statistical Methods, Eighth Edition, Iowa State University Press.

Kolmogorov-Smirnov

tends to be more sensitive near the center of the distribution

Chakravarti, Laha, and Roy, (1967). Handbook of Methods of Applied Statistics, Volume I, John Wiley and Sons, pp. 392-394.

Anderson-Darling

gives more weight to the tails

Stephens, M. A. (1974). *EDF Statistics for Goodness of Fit and Some Comparisons*, Journal of the American Statistical Association, Vol. 69, pp. 730-737.

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www.palisade.com

www.minitab.com

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If $X_1, X_2, X_3, \dots, X_n$ is a random sample from the population, with sample mean \overline{X} , then the limiting form of

$$Z = \frac{\overline{X} - \mu_X}{\sigma / \sqrt{n}}$$

as $n \to \infty~$ is the standard normal distribution

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When there are many variables affecting the outcome, The final result will be normally distributed even if the individual variable distributions are not.







Sums: $Z = X_1 \pm X_2 \pm X_3 \pm X_4 \pm \dots + X_n$

 $Z \rightarrow Normal \mbox{ as } n \mbox{ increases}$

Products: $Z = X_1 \cdot X_2 \cdot X_3 \cdot X_4 \cdot \dots \cdot X_n$

 $Z \rightarrow LogNormal as n increases$

Normal and LogNormal distributions are often employed for analysis even though the underlying population distribution is unknown.



- All variables are random and can be characterized by a statistical distribution with a mean and variance.
- The final result will be normally distributed even if the individual variable distributions are not.



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Analysis Methods

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- Characterizing Variability
- Case Studies
- FatigueCalculator.com
- GlyphWorks

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Response function

Z(X) = Z(Q, L, b, h)

Limit State function

$$g = Z(X) - Z_o = 0$$

Same as the failure function







Need about 10^5 simulations for P(Failure) = 10^{-4}

















Approximate the integral of the joint probability distribution over the failure region





Response Surface Methodology

Response surface methodology is a well established collection of mathematical and statistical techniques for applications where the response of interest is influenced by several variables.

$$Z(X) = f(x_1, x_2, x_3, \cdots , x_n) + \varepsilon$$



Mathematical Representation

$$\begin{split} Z(X) &= A_{_0} + A_{_1}x_{_1} + A_{_2}x_{_2} + A_{_3}x_{_3} + \cdots \cdots \quad \text{Linear} \\ &= B_{_1}x_{_1}^{^2} + B_{_2}x_{_2}^{^2} + B_{_3}x_{_3}^{^2} + \cdots \cdots \quad \text{Incomplete Quadradic} \\ &= C_{_1}x_{_1}x_{_2} + C_{_2}x_{_1}x_{_3} + C_{_3}x_{_2}x_{_3} + \cdots \quad \text{Complete Quadradic} \end{split}$$

	Solutions
Linear	N + 1
Incomplete Quadradic	2N + 1
Complete Quadradic	N(N + 1)/2

Evaluation of Response Surface

Suppose stress is affected by speed and temperature

Factorial Design



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Limit states define probability problems

For example:

 $g = Strength(x_1) - Stress(x_2)$

Prob(g<= 0) = Probability of failure = P_f

Analysis focuses on g = 0













Sometimes called reliability index

 $\boldsymbol{\beta}$ is a number measured in standard deviations

Unlike safety factors, failure probability is directly related to the safety index

$$\mathsf{P}_{\mathsf{f}} = \Phi(-\beta)$$



Requires an efficient numerical search to find the tangent point of a hypersphere (β -sphere) and the limit state function in **u** space











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Standard deviation sensitivity

$$S_{\sigma i} = \frac{\partial P/P}{\partial \sigma_i / \sigma_i}$$

Mean deviation sensitivity

$$\mathsf{S}_{\mu i} = \frac{\partial \mathsf{P} / \mathsf{P}}{\partial \mu_i / \mu_i}$$

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Software General Purpose

Durability

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Variability

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Variability: Every apple on a tree has a different mass.

Uncertainty: The variety of the apple is unknown.

Variability: Fracture toughness of a material

Uncertainty: The correct stress intensity factor solution







COV and LogNormal Distributions

	Standard Deviation, Inx		
001/	1	2	3
COV _x	68.3%	95.4%	99.7%
0.05	1.05	1.11	1.16
0.1	1.10	1.23	1.33
0.25	1.28	1.66	2.04
0.5	1.60	2.64	3.92
1	2.30	5.53	11.1

99.7% of the data is within a factor of \pm 1.33 of the mean for a COV = 0.1



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Barter, S. A., Sharp, P. K., Holden, G. & Clark, G. "Initiation and early growth of fatigue cracks in an aerospace aluminium alloy", *Fatigue & Fracture of Engineering Materials & Structures* **25** (2), 111-125.



7010-T7651

Pre-corroded specimens

300 specimens

246 failed from pits

Crawford et.al."The EIFS Distribution for Anodized and Pre-corroded 7010-T7651 under Constant Amplitude Loading" Fatigue and Fracture of Engineering Materials and Structures, Vol. 28, No. 9 2005, 795-808

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Metals Handbook, 8th Edition, Vol. 1, p64

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Fracture Toughness Mar-M 250 Steel 99.9 % 26 Data Points Median 76.7 99 % **Cumulative Probability** COV 0.06 90 % 50 % 100 70 90 60 80 K_{Ic}, Ksi√in 10 % 1 % 0.1 %

Kies, J.A., Smith, H.L., Romine, H.E. and Bernstein, H, "Fracture Testing of Weldments", ASTM STP 381, 1965, 328-356 Probabilistic Fatigue © 2003-2005 Darrell Socie, University of Illinois at Urbana-Champaign, All Rights Reserved 189 of 352







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ΔS	X	COV
440	14,000	0.12
315	25,000	0.38
280	220,000	0.70
245	1,200,000	0.67
210	12,000,000	1.39

Variability in Fatigue Strength $\frac{\Delta S}{2} = S'_{f} (N_{f})^{b} \quad b \approx -0.085$ COV $C = \sqrt{\prod_{i=1}^{n} (1+C_{x_{i}}^{2})^{a_{i}^{2}} - 1}$

$$C_{s_{f}} = \sqrt{(1+1.39^{2})^{(-.085)^{2}} - 1} = 0.088$$









$$\frac{\Delta \varepsilon}{2} = \frac{\sigma_{f}^{'}(L, \mu_{\sigma_{f}}, \sigma_{\sigma_{f}})}{E} (2N_{f})^{b(N, \mu_{b}, \sigma_{b})} + \varepsilon_{f}^{'}(L, \mu_{\varepsilon_{f}}, \sigma_{\varepsilon_{f}})(2N_{f})^{c(N, \mu_{b}, \sigma_{b})}$$



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Generating Correlated Data

$$z_{1} = \Phi(rand()) \qquad z_{1} = N(0,1)$$

$$z_{2} = \Phi(rand())$$

$$z_{3} = z_{1}\rho + z_{2}\sqrt{1-\rho^{2}}$$

$$\sigma_{f}^{'} = \exp(\mu_{\ln\sigma_{f}^{'}} + \sigma_{\ln\sigma_{f}^{'}}z_{1})$$

$$b = \mu_{b} + \sigma_{b}z_{3}$$







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Virkler, Hillberry and Goel, "The Statistical Nature of Fatigue Crack Propagation", Journal of Engineering Materials and Technology, Vol. 101, 1979, 148-153



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Beware of Correlated Variables $N_{f} = \frac{a_{f}^{1-m/2} - a_{i}^{1-m/2}}{C\Delta S^{m} \pi^{\frac{m}{2}} (1-m/2)}$

 $C \Delta S^m \pi^2 (1-m/2)$

 $\ensuremath{N_{\text{f}}}$ and C are linearly related and should have the same variability, but

$$COV_{N_f} = 0.07$$

 $COV_C = 0.44$

because C and m are correlated.





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Modeling Uncertainty

Analysis Uncertainty $C_U = ?$

The variability in reproducing the original strain life data from the material constants is $C_M \sim 0.44$

COV
$$C = \sqrt{\prod_{i=1}^{n} (1 + C_{X_i}^2)^{a_i^2} - 1}$$

 $1 + C_U^2 = \frac{1 + C_{N_f}^2}{1 + C_M^2}$

C_U = 1.09

90% of the time the analysis is within a factor of 3 ! 99% of the time the analysis is within a factor of 10 !

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Variability from Multiple Sources

$$COV \quad C = \sqrt{\prod_{i=1}^{n} \left(1 + C_{X_i}^{2}\right)^{a_i^2} - 1}$$

Suppose we have 4 variables each with a COV = 0.1

The combined variability is COV = 0.29

Suppose we reduce the variability of one of the variables to 0.05

The combined variability is now COV = 0.27

If all of the COV's are the same, it doesn't do any good to reduce only one of them, you must reduce all of them !

Variability from Multiple Sources

$$COV \ C = \sqrt{\prod_{i=1}^{n} \left(1 + C_{X_i}^{2}\right)^{a_i^2} - 1}$$

Suppose we have 3 variables each with a COV = 0.1 and one with COV = 0.4

The combined variability is COV = 0.65

Suppose we reduce the variability of these variables to 0.05

The combined variability is now COV = 0.60

If one of the COV's is large, it doesn't do any good to reduce the others, you must reduce the largest one !

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Variability Summary Source COV Service Loading Environment Materials Manufacturing Surface Finish 0.5 Stress 0.3 0.1 Strength 0.1 0.1 _ _ _ **_ _ _ _ _ _ _ _ _** Fatigue Lives 1.0 Analysis Uncertainty 1.0 Fatigue life $\propto \left(\frac{\text{Strength}}{\text{Stress}}\right)^5$



Variability: Every apple on a tree has a different mass. Uncertainty: The variety of the apple is unknown.

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Variability: Multiple samples of the same material Uncertainty: What is the material









Fatigue Strength Coefficient

	Variability	Uncertainty	Combined
All Steels	0.12	0.48	0.75
Structural Steel	0.12	0.12	0.24

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At my last seminar everyone hit a golf ball and we recorded the maximum acceleration.

What is the expected variability ?



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Probabilistic Aspects of Fatigue

Case Studies

Professor Darrell F. Socie Department of Mechanical and Industrial Engineering

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- Basic Probability and Statistics
- Statistical Techniques
- Analysis Methods
- Characterizing Variability
- Case Studies
- FatigueCalculator.com
- GlyphWorks

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ASTM Symposium on Probabilistic Aspects of Life Prediction Miami Beach, Florida November 6-7, 2002

R. C. McClung, M. P. Enright, H. R. Millwater^{*}, G. R. Leverant, and S. J. Hudak, Jr. Southwest Research

Slides 6 – 27 used with permission of of Craig McClung

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Anomalies in titanium engine disks

Hard Alpha Very rare Can cause failure Not addressed by safe life methods Enhanced life management process Requested by FAA Developed by engine industry Probabilistic damage tolerance methods Supplement to safe life approach



SwRI and engine industry developed DARWIN with FAA funding







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    Probabilistic Fatigue
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Anomaly Distribution

 # of anomalies per volume of material as function of defect size Library of default anomaly distributions for HA (developed by RISC)



Probability of Detection Curves

Define probability of NDE flaw detection as function of flaw size Can specify different PODs for different zones, schedules Built-in POD library or user-defined POD



Random Inspection Time

"Opportunity Inspections" during on-condition maintenance Inspection time modeled with Normal distribution or CDF table





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Identify regions of component with highest risk









- Anomaly size (initial crack size)
- FCG properties (life scatter)
- Mission histories (stress scatter)

Hard Alpha Defects in Titanium

Initial DARWIN focus on Hard Alpha

Small brittle zone in microstructure

Alpha phase stabilized by N accidentally introduced during melting

Cracks initiate quickly

Extensive industry effort to develop HA distribution

Probabilistic Fatigue









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Fatigue Design and Reliability ESIS Publication 23

J-J. Thomas, G. Perroud, A. Bignonnet and D. Monnet

PSA Peugeot Citroën



Fatigue design at PSA is done with a probabilistic approach that includes analysis of customer usage, production scatter, definition of the appropriate design loads and an acceptance testing criterion.





Variability in loading has two components, how it is used and how it is driven.

Car Usage Highway, city, fully loaded, empty etc.

Driving Style passive, aggressive etc.

The usage of a car is independent of the owners driving style so that the distributions of car usage and driving style can be obtained separately.



Customer surveys

k	1		$c_k \%$	r _{kl} %
1		Unloaded	27	
	1	Highway		10
	2	Good Road		25
	3	Mountain		40
	4	City		25
2		Half Load	58	
	1	Highway		5
	2	Good Road		30
	3	Mountain		30
	4	City		35
3		Fully Loaded	15	
	1	Highway		15
	2	Good Road		25
	3	Mountain		40
	4	City		20

12 Customer Usage Categories

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Extensive field testing for each customer usage category produces a large number of histograms.

Let the usage histogram be denoted U_{kl}



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Thousands of virtual customers can now be generated by combining customer usage with driving style.





Initial design is done on the basis of a single constant amplitude load, $\rm F_{eq}$

Find a constant amplitude load and number of cycles that will produce the same fatigue damage as the customer operating a car for the design life.











Some Statistics

Z = S - F

Suppose we want a probability of failure of 1 in 50,000

$$P_{f} = 2 \times 10^{-5}$$

$$\Phi^{-1}(P_{f}) = 4.1 = \frac{\mu_{Z}}{\sigma_{Z}} = \frac{\mu_{S} - \mu_{F}}{\sqrt{\sigma_{S}^{2} + \sigma_{F}^{2}}}$$

$$F_{n} = \mu_{S}(1 - \beta \text{ COV}_{S})$$

$$F_{n} = \mu_{F}(1 + \alpha \text{ COV}_{F})$$



$$\Phi^{-1}(P_f) = 4.1 = \frac{\mu_Z}{\sigma_Z} = \frac{\mu_S - \mu_F}{\sqrt{\sigma_S^2 + \sigma_F^2}}$$

$$\Phi^{-1}(\mathsf{P}_{\mathsf{f}}) = \frac{\frac{\mu_{\mathsf{S}}}{\mathsf{F}_{\mathsf{n}}} - \frac{1}{1 + \alpha \operatorname{COV}_{\mathsf{F}}}}{\sqrt{\left(\frac{\mu_{\mathsf{S}}}{\mathsf{F}_{\mathsf{n}}} \operatorname{COV}_{\mathsf{S}}\right)^{2} + \left(\frac{\operatorname{COV}_{\mathsf{F}}}{1 + \alpha \operatorname{COV}_{\mathsf{F}}}\right)^{2}}}$$

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Component tests are done with small sample sizes

 $\begin{array}{ll} \text{Confidence limits} & \displaystyle \frac{(N\!-\!1){s_x}^2}{\chi^2_{1\!-\!\alpha,N\!-\!1}} \leq {\sigma_x}^2 \end{array}$

$$\Phi^{-1}(\mathsf{P}_{\mathsf{f}}) = \frac{\frac{\mu_{\mathsf{S}}}{\mathsf{F}_{\mathsf{n}}} - \frac{1}{1 + \alpha \operatorname{COV}_{\mathsf{F}}}}{\sqrt{\left(\frac{\mu_{\mathsf{S}}}{\mathsf{F}_{\mathsf{n}}} \operatorname{COV}_{\mathsf{S}}\right)^{2} \frac{\mathsf{N} - 1}{\chi^{2}_{1-\alpha,\mathsf{N}-1}} + \left(\frac{\operatorname{COV}_{\mathsf{F}}}{1 + \alpha \operatorname{COV}_{\mathsf{F}}}\right)^{2}}}$$





Full scale vehicle simulation done at the end for design final validation

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How does this process work?





extrapolate to more users



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Probabilistic Fatigue

- Given a rainflow histogram for a single user, extrapolate to longer times
- Given rainflow histograms for multiple users, extropolate to more users

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Probabilistic Fatigue

- In the first problem the number of cycles is known but the variability is unknown and must be estimated
- In the second problem the variability is known but the number and location of cycles is unknown and must be estimated

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Airplane Exceedance





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Probabilistic Aspects of Fatigue

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COV and LogNormal Distributions

	Sta	ndard Deviati	on, Inx
001/	1	2	3
COV _X	68.3%	95.4%	99.7%
0.05	1.05	1.11	1.16
0.1	1.10	1.23	1.33
0.25	1.28	1.66	2.04
0.5	1.60	2.64	3.92
1	2.30	5.53	11.1

99.7% of the data is within a factor of \pm 1.33 of the mean for a COV = 0.1

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emer N(0.001, 0.2)	-2.22	0.621	9.92*10*	0.19
e _{min} N(-0.001, 0.2)	-13	0.364	-0.00103	0.18
Material Properties	-6.99	0.621		· · · ·
K'	0.666	0	1440	0
st.	-0.14	0	0.283	0
E	-1.57	0	2.07*105	0
ь	-1.04	0	-0.118	0
¢	-8.02	0	-0.412	0
Sį L(883, 0.1)	0.944	0.132	885	0.091
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Ten Simula	ations		
	Life	COV	
	6470	0.959	
	6930	0.898	
	6710	0.688	
	6640	0.908	
	6580	0.869	
	6470	0.959	
	7010	0.723	
	6690	0.908	
	6170	0.791	
	6560	0.971	
Mean	6623	0.8674	
COV	0.038	0.114	







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e _{min} N(-0.001, 0.2.)	-13	0.359	-0.00101	0.2
Material Properties	-6.99	0.635		-
к	0.666	0	1440	(
n'	-0.14	0	0.283	0
E	-1.57	0	2.07*105	0
b N(-0.118, .25.), CC=83	-1.04	0.2	-0.119	0.3
c N(-0.412, 23)	-8.02	0.506	-0.41	0.0
SJ L(883, 25)	0.944	0.325	861	0.2
Ej L(0.16, 1.15)	2.17	0.0344	0.16	0.01
Stress Concentrators	-4.43	0.305		
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Probabilistic Aspects of Fatigue

GlyphWorks

Professor Darrell F. Socie Department of Mechanical and Industrial Engineering

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- Introduction
- Basic Probability and Statistics
- Statistical Techniques
- Analysis Methods
- Characterizing Variability
- Case Studies
- FatigueCalculator.com
- GlyphWorks

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Probabilistic Fatigue

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Material Data					
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Stp	1000	Log-Normal	0.1		Fatigue strength coefficient
ь	-0.1	None			Fatigue strength exponent
elp	1	None	0.2		Fatigue ductility coefficient
c	-0.5	None			Fatigue ductility exponent
np	0.2	None			Cyclic strain hardening exponent
Кр	1200	LogNomal	0.1		Cyclic strength coefficient
Analysis Properties					
NumCases	10				Number of simulations to run
Damage Sum	1	Loo Nom	5		Uncertainty

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³⁴⁰ of 352





Probabilistic Sensitivity, Channel 6



StatOutput.xls (continued)

Channel # 6

Life Distribution

Sensitivity

1

Total Channels

		Variable	Probabilistic	Deterministic
NumCases	11	Load History Variables	0.89	-5.42
Median	45634	ScaleFactor	0.89	-5.42
COV	1.91	Offset	0.00	0.00
		Stress Concentrators	0.05	-5.51
Probability (%)	Life	Kf	0.05	-5.51
99	817179	Material Properties	0.25	-14.54
90	223648	E	0.00	-4.37
50	45634	Sfp	0.23	3.42
10	9311	b	0.00	-4.39
1	2548	efp	0.00	1.42
		С	0.00	-9.86
		np	0.00	-2.32
		Кр	0.10	1.56
		Analysis Variables	0.37	1.00
		Uncertainty	0.37	1.00

Probabilistic Fatigue

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		Inputs		Outputs		
			Scale		Scale	
Variable	Distribution	Value	Parameter	Value	Parameter	
ScaleFactor	Normal	100	0.25	103	0.18	
Offset	Normal	0	0.10	-0.01	0.08	
Kf	Uniform	3	0.05	3.0	0.04	
E	None	208000	0.00	208000	0.00	
Sfp	Log-Normal	1000	0.10	1054	0.11	
b	None	-0.1	0.00	-0.10	0.00	
efp	None	1	0.20	1.0	0.00	
С	None	-0.5	0.00	-0.50	0.00	
np	None	0.2	0.00	0.20	0.00	
Kp	Log-Normal	1200	0.10	1194	0.10	
Uncertainty	Log-Normal	1	0.50	0.93	0.48	

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StaStress
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Median Life: 2.368E+08 Repeats
Scale Parameter Correlation Coefficient Description
0.1 0 Scale Factor
0.1 Officet
0.05 0 Fatigue Concentration Factor
0.1 Ultimate Tensile Strength
0.1 0 Stress Bange Intercept
0 Main S-N Slope
Number of simulations to run
0.1 0 Analysis Uncertainty
0.1 0.1 0.0 0.1 0.1 0.1 0.1

Probabilistic Fatigue

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Number of Simulations 1			

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    Probabilistic Fatigue
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    Probabilistic Fatigue
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Percentile Extrapolation

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MH_pass7_acc >>	MH_pass4_acc	1	cyh	nCode histog				
MH_pass8_acc	MH_pass5_acc	1	cyh	nCode histog				
MH_pass9_acc <	MH_pass6_acc	1	cyh	nCode histog				
MH_pass10_acc	MH_pass7_acc	1	cyh	nCode histog 🚽				
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Extrapolation Factor (0.5 - 1, but not 1) 0.99	Advanced						
Number of Simulations 1								
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Probabilistic Aspects of Fatigue

